

Formulae for Statistics for Science, Engineering and Technology

A. Descriptive Statistics (for Samples)

A.1 Percentiles: The position of the 100α -th percentile is $R_\alpha = \alpha(1+n) = i+d$, so that the 100α -th percentile is $(1-d)y_i + dy_{i+1}$, where $\alpha = 0.01, 0.02, \dots, 0.99$, i is the largest integer not exceeding R_α and y_i is i -th value of the sorted sample.

A.2 Average and variance are

$$\bar{x} \equiv \frac{1}{n} \sum x; \quad \text{CSS} \equiv \sum (x - \bar{x})^2 = \sum x^2 - \frac{1}{n} (\sum x)^2; \quad s^2 \equiv \frac{\text{CSS}}{n-1}.$$

A.3 Mean and the variance for grouped data:

$$\bar{x} \equiv \frac{1}{n} \sum xf; \quad \text{CSS} \equiv \sum x^2 f - \frac{1}{n} (\sum xf)^2; \quad s^2 \equiv \frac{\text{CSS}}{n-1}.$$

where x 's are the mid values of each class and the sum is over the number of classes.

A.4 Standardized score of an observation x is $z(x) \equiv (x - \bar{x}) / s$.

A.5 Coefficient of Variation : $CV \equiv s / \bar{x}$.

A.6 Coefficient of Skewness : $CS \equiv \frac{3(\bar{x} - \tilde{x})}{s}$, where \tilde{x} is the sample median.

B. Glossary of Probability of Set Events (Two Sets A and B)

B.1 A but not B (=Only A): $P(AB') = P(A) - P(AB)$.

B.2 $P(A'B') + P(AB') + P(A'B) + P(AB) = 1$

B.3 Law of Addition

B.3a $P(A \cup B) = P(AB') + P(A'B) + P(AB)$.

B.3b $P(A \cup B) = P(A) + P(B) - P(AB)$.

B.3c $P(A \cup B) = 1 - P(A \cup B)' = 1 - P(A'B')$. De Morgan's Law

B.4 $P(AB)' = P(A' \cup B')$. De Morgan's Law

B.5 Law of Multiplication

$$P(A|B) = \frac{P(AB)}{P(B)}, \quad P(B) \neq 0; \quad P(AB) = P(A)P(B|A) = P(B)P(A|B)$$

B.6 Independence: $P(A | B') = P(A) = P(A | B)$, $P(AB) = P(A)P(B)$. .

For intersection of three sets A, B and C , the following results are important:

B.7 None: $P(A'B'C') = P(A')P(B')P(C')$ iff A, B and C are independent.

B.8 $P(A \cup B \cup C) = P(A) + P(B) + P(C) - [P(AB) + P(AC) + P(BC)] + P(ABC)$.

C. Discrete Probability Distributions

C.1 $P(a \leq X \leq b) = \sum_x f(x)$; $P(X \leq b) = \sum_{x \leq b} f(x)$,

C.2 $\mu \equiv E(X) = \sum_x x f(x)$.

C.3 $E(X^2) = \sum x^2 f(x)$, $\sigma^2 \equiv E(X - \mu)^2 = E(X^2) - \mu^2$.

C.4 The Binomial Distribution $B(n, p)$: $f(x) = \binom{n}{x} p^x q^{n-x}$; $x = 0, 1, \dots, n$; $0 < p < 1$;

$q = 1 - p$; $\mu = np$, $\sigma^2 = npq$.

C.5 The Geometric Distribution: $f(x) = q^x p$, $x = 0, 1, 2, \dots$; $q = 1 - p$; $\mu = 1/p$, $\sigma^2 = q/p^2$.

C.7 The Hypergeometric Distribution

$f(x) = \binom{K}{x} \binom{N-K}{n-x} \div \binom{N}{n}$, $\max\{0, n - (N - K)\} \leq y \leq \min\{n, K\}$; $\mu = np$,

$\sigma^2 = (1-c) npq$, $(N-1) c = n-1$, $p = (K/N)$, $q = 1 - p$.

The mass function of hypergeometric distribution can also be written as

$P(X = x) = \binom{n}{x} P(D_1 D_2 \dots D_x D'_{x+1} \dots D'_n) = \binom{n}{x} (P_x^K P_{n-x}^{N-K} \div P_n^N)$.

C.8 The Poisson Distribution $f(x) = \frac{(\lambda t)^x}{x!} e^{-\lambda t}$; $x = 0, 1, \dots$; $\mu = \lambda t$, $\sigma^2 = \lambda t$.

D. Continuous Probability Distributions

For a continuous random variable X with pdf $f(x)$,

$$\text{D.1 } P(a < X < b) = \int_a^b f(x)dx; \quad F(k) = \int_{-\infty}^k f(x)dx \text{ where } k \text{ is a particular value of } x.$$

$$\text{D.2 } \mu \equiv E(X) = \int_{-\infty}^{\infty} x f(x)dx.$$

$$\text{D.3 } E(X^2) = \int_{-\infty}^{\infty} x^2 f(x)dx \quad \sigma^2 \equiv V(X) = E(X^2) - \mu^2.$$

D.4 The Normal Distribution $N(\mu, \sigma^2)$

D.5 The Exponential Distribution: $f(x) = \frac{1}{\beta} e^{-x/\beta}$, $0 \leq x$; $\mu = \beta$, $\sigma^2 = \beta^2$.

D.6 Waiting Time Distribution: $f(t) = \lambda e^{-\lambda t}$, $0 \leq t$; $\mu = 1/\lambda$, $\sigma^2 = 1/\lambda^2$,

D.7 The Gamma Distribution: $f(x) = \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-x/\beta}$, $0 < x$; $0 < \alpha$, $0 < \beta$, $\mu = \alpha\beta$,
 $\sigma^2 = \alpha\beta^2$.

E. Sampling Distributions

E.1 Reproductive Theorem: Suppose that X has a distribution with mean μ and variance σ^2 .

Additionally if the distribution is normal, then $\frac{\sum X - n\mu}{\sqrt{n\sigma^2}} = \frac{\bar{X} - \mu}{\sqrt{\sigma^2/n}} = Z$.

E.2 Suppose that Y has a distribution with mean μ and variance σ^2 . However if the distribution is not normal but $30 \leq n$, then

$$\frac{\sum X - n\mu}{\sqrt{n\sigma^2}} = \frac{\bar{X} - \mu}{\sqrt{\sigma^2/n}} \approx Z \text{ (weak)}.$$

This is known as Central Limit Theorem (CLT). In case σ^2 is unknown,

$$\frac{\sum X - n\mu}{\sqrt{nS^2}} = \frac{\bar{X} - \mu}{\sqrt{S^2/n}} \approx Z \text{ (weaker)}$$

by Slutsky's Theorem (CT). Note S^2 converges stochastically to σ^2 .

E.3 The Student T- statistic is defined by $T = \frac{\bar{X} - \mu}{\sqrt{S^2 / n}}$, with $\nu = n - 1$

E.4 The Sampling Distribution of the Proportion

$$\frac{X - np}{\sqrt{npq}} = \frac{(X/n) - p}{\sqrt{pq/n}} \approx Z.$$

E.5 The binomial event $X = x$ can be inflated to $x - 0.50 \leq X \leq x + 0.50$ if $4 \leq np$ and $4 \leq nq$.

The $100(1 - \alpha)\%$ percentiles of a random variable Z is denoted by $z_{1-\alpha}$ or often by z_α . For example 95th percentile of the standard normal distribution is given by $z_{0.95} \approx 1.645$.

F. Statistical Estimation (with Random Sample / Samples)

F.1 Confidence Interval Estimates of the Mean μ

F.1.1 CI for μ , (σ known, any n , normal): $\bar{x} \mp z_{\alpha/2} \left(\frac{\sigma}{\sqrt{n}} \right)$.

F.1.2 ACI for μ , (σ known, large n , nonnormal): $\bar{x} \mp z_{\alpha/2} \left(\frac{\sigma}{\sqrt{n}} \right)$.

F.1.2a sample size for estimating μ : $n = \frac{z_{\alpha/2}^2 \sigma^2}{e^2}$, where $P(|\bar{X} - \mu| \leq e) = 1 - \alpha$.

F.1.3 ACI for μ , (σ unknown, large n , nonnormal): $\bar{x} \mp z_{\alpha/2} \left(\frac{s}{\sqrt{n}} \right)$.

F.1.3a sample size for estimating μ : $n = \frac{z_{\alpha/2}^2 s^2}{e^2}$, where $P(|\bar{X} - \mu| \leq e) = 1 - \alpha$.

F.1.4 CI for μ , (σ unknown, $n \geq 2$, normal): $\bar{x} \mp t_{\alpha/2} \left(\frac{s}{\sqrt{n}} \right)$ for any $n \geq 2$.

F.2 Confidence Interval for $\mu_1 - \mu_2$ (Based on Random and Independent Samples)

F.2.1 CI for $\mu_1 - \mu_2$, (σ_i^2 known, any n_i , normal):

$$(\bar{x}_1 - \bar{x}_2) \mp z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}.$$

F.2.2 ACI for $\mu_1 - \mu_2$, (σ_i^2 known, large n_i , nonnormal):

$$(\bar{x}_1 - \bar{x}_2) \mp z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}.$$

F.2.3 CI for $\mu_1 - \mu_2$, (σ_i^2 unknown, large n_i , nonnormal):

$$(\bar{x}_1 - \bar{x}_2) \mp z_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}.$$

F.2.4 CI for $\mu_1 - \mu_2$, (small n_i , unknown $\sigma_1^2 = \sigma_2^2$ but equal, normal):

$$(\bar{x}_1 - \bar{x}_2) \mp t_{\alpha/2} \sqrt{\frac{s_w^2}{n_1} + \frac{s_w^2}{n_2}}, \quad s_w^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{(n_1 - 1) + (n_2 - 1)}, \quad \nu = (n_1 - 1) + (n_2 - 1).$$

F.3 Confidence Interval for Proportion p

F.3.1 CI for p when n large: $\hat{p} \mp z_{\alpha/2} \sqrt{\hat{p}\hat{q}/n}$.

F.3.1 a Large sample size for estimating p : $n = \frac{z_{\alpha/2}^2}{e^2} \hat{p}\hat{q}$ where e is the error in estimation.

F.3.2 CI for $p_1 - p_2$ with large sample sizes :

$$\hat{p}_1 - \hat{p}_2 \mp z_{\alpha/2} \sqrt{\frac{\hat{p}_1\hat{q}_1}{n_1} + \frac{\hat{p}_2\hat{q}_2}{n_2}}.$$

G. Testing of Hypotheses (with Random Sample/ Samples)

Reject H_0 for $p\text{-value} \leq \alpha$; Don't reject H_0 for $0 \leq \alpha < p\text{-value}$.

G.1 Testing of a Mean μ

G.1.1 σ known, normal: $z = \frac{\bar{x} - \mu_0}{\sqrt{\sigma^2/n}}$.

G.1.2 σ known, large n , nonnormal: $z \approx \frac{\bar{x} - \mu_0}{\sqrt{\sigma^2/n}}$.

G.1.3 σ unknown, large sample: $z \approx \frac{\bar{y} - \mu_0}{\sqrt{s^2/n}}$

G.1.4 σ unknown, normal population : $t = \frac{\bar{x} - \mu_0}{\sqrt{s^2/n}}$, ($\nu = n - 1 \geq 1$).

G.2 Testing $\mu_1 - \mu_2 = \delta$ (Random and Independent Samples)

G.2.1 known σ_i , large n_i , normal: $z = \frac{\bar{x}_1 - \bar{x}_2 - \delta_0}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$.

G.2.2 known σ_i , large n_i , nonnormal: $z \approx \frac{\bar{x}_1 - \bar{x}_2 - \delta_0}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$.

G.2.3 unknown σ_i , large n_i , nonnormal: $z \approx \frac{\bar{x}_1 - \bar{x}_2 - \delta_0}{\sqrt{(s_1^2/n_1) + (s_2^2/n_2)}}$.

G.2.3 Small n_i , unknown $\sigma_1^2 = \sigma_2^2$, normal)

$$t = \frac{\bar{x}_1 - \bar{x}_2 - \delta_0}{\sqrt{(s_w^2/n_1) + (s_w^2/n_2)}}, \quad s_w^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}, \quad \nu = (n_1 - 1) + (n_2 - 1),$$

where s_w^2 is the weighted or pooled combined variance $\min(s_1^2, s_2^2) \leq s_w^2 \leq \max(s_1^2, s_2^2)$.

G.3.1 Testing of a proportion

Large sample: $z \approx \frac{\hat{p} - p_0}{\sqrt{p_0 q_0 / n}}$, where $q_0 = 1 - p_0$.

H. Linear Regression Analysis (Degrees of Freedom: $\nu = n - 2$)

H.1 Line of Best Fit

H.1.0 Assumed Model: $y = \beta_0 + \beta_1 x + \varepsilon$ for a given x , $Y \sim N(\mu(x), \sigma^2)$ where $\mu(x) = \beta_0 + \beta_1 x$.

Estimated Model: $\hat{\mu}(x) = \hat{\beta}_0 + \hat{\beta}_1 x$ for a given x . $\hat{\mu}(x)$ is often denoted by \hat{y} . Error $\varepsilon(x) = y - \mu(x)$ is estimated by $e = y - \hat{\mu}(x)$. Also $\Delta\mu(x) = \mu(x+h) - \mu(x) = \beta_1 h$.

H.1.1 $\hat{\beta}_1 = \frac{s_{xy}}{s_{xx}}$, $s_{xy} = \sum xy - \frac{1}{n}(\sum x)(\sum y)$, $s_{xx} = \sum x^2 - \frac{1}{n}(\sum x)^2$, $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$.

H.1.2 Coefficient of correlation: $r = \frac{s_{xy}}{\sqrt{s_{xx} s_{yy}}}$. ($r\sqrt{s_{yy}} = \hat{\beta}_1\sqrt{s_{xx}}$)

H.1.3 $CSS = \sum y^2 - \frac{1}{n}(\sum y)^2$, $SSR = \hat{\beta}_1 s_{xy}$, $SSE = CSS - SSR$.

H.1.4 Estimate of σ^2 : $s_e^2 = SSE / (n - 2)$ often denoted by MSE .

H.1.5 $R^2 = \frac{\text{Explained variation}}{\text{Total variation}} = \frac{SSR}{CSS}$.

H.2 Inference Regarding the Regression Coefficients

H.2.1 $100(1-\alpha)\%$ confidence interval for β_0 : $\hat{\beta}_0 \mp t_{\alpha/2} \sqrt{\left(\frac{1}{n} + \frac{\bar{x}^2}{s_{xx}}\right) MSE}$.

H.2.2 Testing the hypothesis $H_0: \beta_0 = c$: $t = \frac{\hat{\beta}_0 - c}{\sqrt{\left(\frac{1}{n} + \frac{\bar{x}^2}{s_{xx}}\right) MSE}}$.

H.2.3 $100(1-\alpha)\%$ confidence interval for β_1 : $\hat{\beta}_1 \mp t_{\alpha/2} \sqrt{\frac{MSE}{s_{xx}}}$.

H.2.4 Testing the hypothesis $H_0: \beta_1 = c$: $t = \frac{\hat{\beta}_1 - c}{\sqrt{MSE / s_{xx}}}$.

H.2.5 Testing the hypothesis $H_0: \rho = 0$: $t = \frac{r\sqrt{n-2}}{\sqrt{1-r^2}}$. ($\rho\sigma_y = \beta_1\sigma_x$)

Inference Regarding the Response Variable

H.3.1 $100(1-\alpha)\%$ Confidence Interval of $\mu(x)$: $\hat{\mu}(x) \pm t_{\alpha/2} \sqrt{\left(\frac{1}{n} + \frac{(x-\bar{x})^2}{s_{xx}}\right) MSE}$.

H.3.2 $100(1-\alpha)\%$ Prediction Interval for an Individual Y for a given x :

$$\hat{\mu}(x) \pm t_{\alpha/2} \sqrt{\left(1 + \frac{1}{n} + \frac{(x-\bar{x})^2}{s_{xx}}\right) MSE}.$$

GOLDEN TABLE (Sigma-Distribution)

	$X \sim N(\mu, \sigma^2)$	$X \not\sim N(\mu, \sigma^2)$
σ known (Moya moujud)	$\frac{\bar{X} - \mu}{\sqrt{\sigma^2 / n}} = Z$	Small n : Assume $X \sim N(\mu, \sigma^2)$; Exact $\frac{\bar{X} - \mu}{\sqrt{\sigma^2 / n}} = Z$. Large n : $\frac{\bar{X} - \mu}{\sqrt{\sigma^2 / n}} \approx Z$ (CLT)
s known (Ma fiyy moya)	$\frac{\bar{X} - \mu}{\sqrt{S^2 / n}} = T$	Large n : $\frac{\bar{X} - \mu}{\sqrt{S^2 / n}} \approx Z$ (ST)