

FORMULAE FOR STAT 319

(Edited by Anwar Joarder)

A. Descriptive Statistics (for Samples)

A.1 Mean and variance are $\bar{y} = \frac{1}{n} \sum y$ and

$$s^2 = \frac{TSS}{n-1} \text{ where } TSS = \sum (y - \bar{y})^2 = \sum y^2 - \frac{1}{n} (\sum y)^2.$$

A.2 Quartiles: $R_\alpha = \alpha \frac{1+n}{4} = i+d$, $\alpha = 1, 2, 3$; $Q_\alpha = (1-d)y_{(i)} + dy_{(i+1)}$.

A.3 Mean and the variance for grouped data:

$$\bar{y} = \frac{1}{n} \sum yf, \quad s^2 = \frac{TSS}{n-1}, \quad TSS = \sum y^2 f - \frac{1}{n} (\sum yf)^2.$$

A.4 Coefficient of Variation: $CV = s/\bar{y}$.

A.5 Coefficient of Skewness: $CS = \frac{\bar{y} - \tilde{y}}{s/3}$.

B. Glossary of Probability of Set Events (Two Sets)

| | Verbal Description of Event | Probability |
|-----|---|---|
| B.1 | <i>A but not B</i> (=Only $A = A$ alone) | $P(A \cap \bar{B}) = P(A) - P(A \cap B)$ |
| B.2 | <i>B but not A</i> (=Only $B = B$ alone) | $P(\bar{A} \cap B) = P(B) - P(A \cap B)$ |
| B.3 | <i>None (=Neither A nor B)</i> | $P(\bar{A} \cap \bar{B}) = P(\overline{A \cup B}) = 1 - P(A \cup B)$ |
| B.4 | <i>Exactly one</i> | $P(A \cap \bar{B}) + P(\bar{A} \cap B)$ |
| B.5 | <i>Both</i> (=Exactly two = Two) | $P(A \cap B) = 1 - P(\overline{A \cap B}) = 1 - P(\bar{A} \cup \bar{B})$ $P(A \cap B) = P(B)P(A B) = P(A)P(B A)$ $P(A \cap B) = P(A)P(B)$ iff A and B are independent |
| B.6 | <i>Not both</i> | $P(\overline{A \cap B}) = 1 - P(A \cap B)$ |
| B.7 | <i>A given B</i> | $P(A B) = \frac{P(A \cap B)}{P(B)}$ if $P(B) \neq 0$ $P(A B) = P(A)$ iff A and B are independent |
| B.8 | <i>At least one of the two</i> (= A or B) | $P(A \cup B) = P(\text{exactly one}) + P(\text{exactly two})$ $= P(A \cap \bar{B}) + P(\bar{A} \cap B) + P(A \cap B)$ $= P(A) + P(B) - P(A \cap B)$ $= 1 - P(\overline{A \cup B}) = 1 - P(\bar{A} \cap \bar{B})$ |

Independence:

$$P(A | \bar{B}) = P(A | B) = P(A)$$

$$P(B | \bar{A}) = P(B | A) = P(B)$$

For intersection of three sets A, B and C , the following results are important:

$$B.9 \quad P(\text{none}) + P(\text{at least one}) = 1$$

$$B.10 \quad P(\text{none}) = P(\bar{A} \cap \bar{B} \cap \bar{C}) \\ = P(\bar{A})P(\bar{B})P(\bar{C}) \text{ by independence}$$

$$B.11 \quad P(\text{at least one}) = P(A \cup B \cup C) \\ = P(A) + P(B) + P(C) - [P(A \cap B) + P(A \cap C) + P(B \cap C)] + P(A \cap B \cap C)$$

C. Discrete Probability Distributions

$$C.0a \quad \mu = E(Y) = \sum yp(y), \quad (p89)$$

$$C.0b \quad E(Y^2) = \sum y^2 p(y), \quad \sigma^2 = E(Y - \mu)^2 = E(Y^2) - \mu^2, \quad (p96)$$

| | Probability Density function $p(x)$ | Mean (μ) and Variance (σ^2) |
|-----|---|---|
| C.1 | The Binomial Distribution: $B(n, p)$ (p119) $f(y) = \binom{n}{y} p^y (1-p)^{n-y}, \quad y = 0, 1, \dots, n,$ where $\binom{n}{y} = \frac{n(n-1)\dots(n-y+1)}{y!}.$ | $\mu = E(Y) = np$ $\sigma^2 = V(Y) = np(1-p).$ |
| C.2 | The Hypergeometric Distribution (p128) $f(y) = \frac{\binom{K}{y} \binom{N-K}{n-y}}{\binom{N}{n}}, \quad y = 0, 1, \dots,$ | $\mu = E(Y) = n (K / N)$ |
| C.3 | The Poisson Distribution (p136) $f(y) = \frac{e^{-\lambda t} (\lambda t)^y}{y!}, \quad y = 0, 1, \dots$ | $\mu = E(Y) = \lambda t$ $\sigma^2 = V(Y) = \lambda t$ |

D. Continuous Probability Distributions

For a continuous random variable Y with pdf $f(y)$

$$D.0 \quad P(a < Y < b) = \int_a^b f(y)dy, \quad P(Y \leq u) = \int_{-\infty}^u f(y)dy$$

$$D.0a \quad \mu = E(Y) = \int_{-\infty}^{\infty} yf(y)dy, \quad (p89)$$

$$D.0b \quad E(Y^2) = \int_{-\infty}^{\infty} y^2 f(y)dx, \quad \sigma^2 = E(Y - \mu)^2 = E(Y^2) - \mu^2, \quad (p96)$$

| | Probability Density function | Mean and Variance |
|-----|--|--|
| D.1 | The Normal Distribution: $N(\mu, \sigma^2)$ | Mean = $E(Y) = \mu$ Variance = $V(Y) = \sigma^2$ |
| D.2 | The Exponential Distribution $f(y) = \frac{1}{\beta} e^{-y/\beta}, \quad 0 < y$ | Mean = $E(Y) = \beta$ Variance = $V(Y) = \beta^2$ |
| D.3 | Waiting Time Distribution $f(t) = \lambda e^{-\lambda t}, \quad 0 < t$ | Mean = $E(T) = 1/\lambda$ Variance = $V(T) = 1/\lambda^2$ |
| D.4 | The Gamma Distribution $f(y) = \frac{1}{\beta^\alpha \Gamma(\alpha)} y^{\alpha-1} e^{-y/\beta}, \quad 0 < y, \quad 0 < \alpha, \quad 0 < \beta$ | Mean = $E(Y) = \alpha\beta$ Variance = $V(Y) = \alpha\beta^2$ |

E. Sampling Distributions

E.1 Suppose that Y has a distribution with mean μ and variance σ^2 . Additionally if

the distribution is normal then $\frac{\sum Y - n\mu}{\sqrt{n\sigma^2}} = \frac{\bar{Y} - \mu}{\sqrt{\sigma^2/n}} = Z$.

E.2 Suppose that Y has a distribution with mean μ and variance σ^2 . However if the

distribution is not normal but $n \geq 30$, then $\frac{\sum Y - n\mu}{\sqrt{nS^2}} = \frac{\bar{Y} - \mu}{\sqrt{S^2/n}} \approx Z$ (p210). This is known as **Central Limit Theorem**. Note that σ^2 is estimated by s^2 (SLLN).

E.3 The Student t statistic is defined by $T = \frac{\bar{Y} - \mu}{S/\sqrt{n}}$, with $\nu = n - 1$ (p220)

E.4 The Sampling Distribution of the Proportion (p258)

$$\frac{Y - n\pi}{\sqrt{n\pi(1-\pi)}} = \frac{Y/n - \pi}{\sqrt{\pi(1-\pi)/n}} \approx Z$$

F. Statistical Estimation (with Random Sample / Samples)

F.1 Confidence Interval Estimates of the Mean μ

F.1.0 CI for μ , (σ known, any n , normal): $\bar{y} \mp z_{\alpha/2}(\sigma/\sqrt{n})$ for (235)

F.1.1 ACI for μ , (σ known, nonnormal, large n): $\bar{y} \mp z_{\alpha/2}(\sigma/\sqrt{n})$ (cf. p235).

F.1.1(a) sample size for estimating μ : $n = z_{\alpha/2}^2 \sigma^2 / d^2$ (p237).

F.1.1 ACI for μ , (σ unknown, large n , nonnormal): $\bar{y} \mp z_{\alpha/2}(s/\sqrt{n})$ (cf. p235).

F.1.3 CI for μ , (σ unknown, $n \geq 2$, normal): $\bar{y} \mp t_{\alpha/2}(s/\sqrt{n})$ for any $n \geq 2$ (p239).

F.2 Confidence Interval for $\mu_1 - \mu_2$ (Based on Random and Independent Samples)

F.2.0 CI for $\mu_1 - \mu_2$, (σ_i^2 known, normal): $(\bar{y}_1 - \bar{y}_2) \mp z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$ (cf. p247)

F.2.1 ACI for $\mu_1 - \mu_2$, (σ_i^2 known, large n_i , nonnormal):

$$(\bar{y}_1 - \bar{y}_2) \mp z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \text{ (cf. p247)}$$

F.2.2 CI for $\mu_1 - \mu_2$, (σ_i^2 unknown, large n_i , nonnormal): $(\bar{y}_1 - \bar{y}_2) \mp z_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$ (p247)

F.2.3 CI for $\mu_1 - \mu_2$, (small n_i , unknown $\sigma_1^2 = \sigma_2^2$ but unknown, normal):

$$(\bar{y}_1 - \bar{y}_2) \mp t_{\alpha/2} \sqrt{\frac{s^2}{n_1} + \frac{s^2}{n_2}}, \quad s^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{(n_1 - 1) + (n_2 - 1)}, \quad \nu = (n_1 - 1) + (n_2 - 1), \text{ (p249)}$$

F.2.4 CI for $\mu_1 - \mu_2$, (small $n_1 = n_2$, $\sigma_1^2 \neq \sigma_2^2$, normal):

$$(\bar{y}_1 - \bar{y}_2) \mp t_{\alpha/2} \sqrt{\frac{s_1^2}{n} + \frac{s_2^2}{n}}, \quad \nu = (n - 1) + (n - 1) \text{ (p249)}$$

F.2.5 ACI for $\mu_1 - \mu_2$, (small $n_1 \neq n_2$, $\sigma_1^2 \neq \sigma_2^2$, normal):

$$(\bar{y}_1 - \bar{y}_2) \mp t_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}, \quad v = \frac{(s_1^2/n_1 + s_2^2/n_2)^2}{(s_1^2/n_1)^2/(n_1-1) + (s_2^2/n_2)^2/(n_2-1)}, \quad (\text{p251})$$

F.2.6 CI for $\mu_1 - \mu_2$ using matched pairs): $\bar{d} \mp t_{\alpha/2} \sqrt{s_d^2/n}$, (df = n - 1), (p254)

F.3 Confidence Interval for Proportion p

F.3.1 CI for π when n large: $\hat{p} \mp z_{\alpha/2} \sqrt{\hat{p}(1-\hat{p})/n}$, (p258)

F.3.1 (a) Large sample size for estimating p : $n = z_{\alpha/2}^2 \hat{p}(1-\hat{p})/e^2$. (p 259)

F.3.2 CI for $\pi_1 - \pi_2$ with large sample sizes :

$$\hat{p}_1 - \hat{p}_2 \mp z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

F.4 Confidence Interval for Variance σ^2 : $[(n-1)s^2/l, (n-1)s^2/u]$

where $l = \chi_{\alpha/2}^2$ and $u = \chi_{1-\alpha/2}^2$

G. Testing of Hypotheses (with Random Sample/ Samples)

Reject H_0 for $\alpha \geq p\text{-value}$; Don't reject H_0 for $\alpha < p\text{-value}$.

G.1 Testing of a Mean μ

G.1.0 σ known, normal: $z = \frac{\bar{y} - \mu_0}{\sqrt{\sigma^2/n}}$ (p300)

G.1.1 σ known, large n , nonnormal: $z \approx \frac{\bar{y} - \mu_0}{\sqrt{\sigma^2/n}}$

G.1.2 σ unknown, large sample: $z \approx \frac{\bar{y} - \mu_0}{\sqrt{s^2/n}}$ (p300)

| H_0 | H_a | RR (for H_0) | p-value |
|---------------|------------------|---|----------------|
| $\mu = \mu_0$ | $\mu < \mu_0$ | $z < -z_\alpha$ | $P(Z < z)$ |
| $\mu = \mu_0$ | $\mu > \mu_0$ | $z > z_\alpha$ | $P(Z > z)$ |
| $\mu = \mu_0$ | $\mu \neq \mu_0$ | $z < -z_{\alpha/2}$ or $z > z_{\alpha/2}$ | $2 P(Z > z)$ |

G.1.3 σ unknown, normal population : $t = \frac{\bar{y} - \mu_0}{\sqrt{s^2/n}}$, $(\nu = n - 1 \geq 1)$ (p304)

| H_0 | H_a | RR (for H_0) | p-value |
|---------------|------------------|---|----------------|
| $\mu = \mu_0$ | $\mu < \mu_0$ | $t < -t_\alpha$ | $P(T < t)$ |
| $\mu = \mu_0$ | $\mu > \mu_0$ | $t > t_\alpha$ | $P(T > t)$ |
| $\mu = \mu_0$ | $\mu \neq \mu_0$ | $t < -t_{\alpha/2}$ or $t > t_{\alpha/2}$ | $2 P(T > t)$ |

G.2 Testing $\mu_1 - \mu_2 = \delta$ (Random and Independent Samples)

G.2.0 known σ_i , any n_i , normal: $z = \frac{\bar{y}_1 - \bar{y}_2 - \delta}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$

G.2.1 known σ_i , large n_i , nonnormal: $z \approx \frac{\bar{y}_1 - \bar{y}_2 - \delta}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$

G.2.2 unknown σ_i , large n_i , nonnormal: $z \approx \frac{\bar{y}_1 - \bar{y}_2 - \delta}{\sqrt{(s^2/n_1) + (s^2/n_2)}}$,

$$s^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2} , \nu = (n_1 - 1) + (n_2 - 1)$$

Table for the above three tests are given below:

| H_0 | H_a | RR (for H_0) | p-value |
|--------------------------|-----------------------------|---|----------------|
| $\mu_1 - \mu_2 = \delta$ | $\mu_1 - \mu_2 < \delta$ | $z < -z_\alpha$ | $P(Z < z)$ |
| $\mu_1 - \mu_2 = \delta$ | $\mu_1 - \mu_2 > \delta$ | $z > z_\alpha$ | $P(Z > z)$ |
| $\mu_1 - \mu_2 = \delta$ | $\mu_1 - \mu_2 \neq \delta$ | $z < -z_{\alpha/2}$ or $z > z_{\alpha/2}$ | $2 P(Z > z)$ |

G.2.3 Small n_i , unknown $\sigma_1^2 = \sigma_2^2$, normal)

$$t = \frac{\bar{y}_1 - \bar{y}_2 - \delta}{\sqrt{(s^2/n_1) + (s^2/n_2)}} , s^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2} , \nu = (n_1 - 1) + (n_2 - 1) , (p308)$$

G.2.4 Small $n_1 = n_2$, $\sigma_1^2 \neq \sigma_2^2$, normal)

$$t = \frac{\bar{y}_1 - \bar{y}_2 - \delta}{\sqrt{(s_1^2/n) + (s_2^2/n)}} , \nu = (n - 1) + (n - 1)$$

G.2.5 Small $n_1 \neq n_2$, $\sigma_1^2 \neq \sigma_2^2$, normal)

$$t \approx \frac{\bar{y}_1 - \bar{y}_2 - \delta}{\sqrt{(s_1^2/n_1) + (s_2^2/n_2)}}, \quad v = \frac{(s_1^2/n_1 + s_2^2/n_2)^2}{(s_1^2/n_1)^2/(n_1-1) + (s_2^2/n_2)^2/(n_2-1)}, \quad (\text{p309}).$$

Tables for the above three tests are provided below:

| H_0 | H_a | RR (for H_0) | p-value |
|--------------------------|-----------------------------|---|----------------|
| $\mu_1 - \mu_2 = \delta$ | $\mu_1 - \mu_2 < \delta$ | $t < -t_\alpha$ | $P(T < t)$ |
| $\mu_1 - \mu_2 = \delta$ | $\mu_1 - \mu_2 > \delta$ | $t > t_\alpha$ | $P(T > t)$ |
| $\mu_1 - \mu_2 = \delta$ | $\mu_1 - \mu_2 \neq \delta$ | $t < -t_{\alpha/2}$ or $t > t_{\alpha/2}$ | $2 P(T > t)$ |

G.2.5 Testing $\mu_1 - \mu_2$ using matched pairs): $t = \bar{d} / \sqrt{s_d^2/n}$, (df = n - 1), (p310)

G.3.1 Testing of a proportion

Large sample: $z \approx \frac{\hat{p} - p}{\sqrt{p(1-p)/n}}$.

| H_0 | H_a | RR (for H_0) | p-value |
|-----------|--------------|---|----------------|
| $p = p_0$ | $p < p_0$ | $z < -z_\alpha$ | $P(Z < z)$ |
| $p = p_0$ | $p > p_0$ | $z > z_\alpha$ | $P(Z > z)$ |
| $p = p_0$ | $p \neq p_0$ | $z < -z_{\alpha/2}$ or $z > z_{\alpha/2}$ | $2 P(Z > z)$ |

G.3.2 Testing the difference of two proportions

Large n_i : $z \approx \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})/n_1 + \hat{p}(1-\hat{p})/n_2}}$ where $\hat{p} = \frac{n_1\hat{p}_1 + n_2\hat{p}_2}{n_1 + n_2} = \frac{y_1 + y_2}{n_1 + n_2}$

| H_0 | H_a | RR (for H_0) | p-value |
|-----------------|----------------|---|----------------|
| $p_1 = p_2$ | $p_1 < p_2$ | $z < -z_\alpha$ | $P(Z < z)$ |
| $\pi_1 = \pi_2$ | $p_1 > p_2$ | $z > z_\alpha$ | $P(Z > z)$ |
| $\pi_1 = \pi_2$ | $p_1 \neq p_2$ | $z < -z_{\alpha/2}$ or $z > z_{\alpha/2}$ | $2 P(Z > z)$ |

G.4.1 Testing a variance $\chi^2 = (n-1)s^2 / \sigma_0^2$

To test $H_0 : \sigma^2 = \sigma_0^2$, $H_0 : \sigma^2 > \sigma_0^2$ use $\chi^2 > \chi_\alpha^2$

To test $H_0 : \sigma^2 = \sigma_0^2$, $H_0 : \sigma^2 < \sigma_0^2$ use $\chi^2 < \chi_\alpha^2$

To test $H_0 : \sigma^2 = \sigma_0^2$, $H_0 : \sigma^2 \neq \sigma_0^2$ use $\chi^2 < \chi_{1-\alpha/2}^2$ or $\chi^2 < \chi_{\alpha/2}^2$

H. Linear Regression Analysis (Degrees of Freedom: $\nu = n - 2$)

H.1 Line of Best Fit

H.1.0 Model : $y = \beta_0 + \beta_1 x + \varepsilon$ for a given x , $Y \sim N(\mu, \sigma^2)$ where $\mu = \beta_0 + \beta_1 x$

Fitted Model: $\hat{y} = \hat{\mu} = \hat{\beta}_0 + \hat{\beta}_1 x$ for a given x

H.1.1 $\hat{\beta}_1 = \frac{s_{xy}}{s_{xx}}$, $s_{xy} = \sum xy - (\sum x)(\sum y)/n$, $s_{xx} = \sum x^2 - (\sum x)^2/n$, $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$.
(p536)

H.1.2 Pearson product moment coefficient of correlation: $r = \frac{s_{xy}}{\sqrt{s_{xx} s_{yy}}}$ (p392)

H.1.3 Standard error of the estimate: $s = \sqrt{SSE/(n-2)} = \sqrt{MSE}$

H.1.4 $TSS = \sum (y - \bar{y})^2 = \sum y^2 - \frac{(\sum y)^2}{n}$, $SSR = b s_{xy} = b^2 s_{xx}$, $SSE = TSS - SSR$

H.1.5 $R^2 = \frac{\text{Explained variation}}{\text{Total variation}} = \frac{SSR}{TSS}$

H.2 Inference Regarding the Regression Coefficient

H.2.1 100(1- α)% confidence interval for β_0 : $\hat{\beta}_0 \mp \sqrt{\left\{ \frac{1}{n} + \frac{\bar{x}^2}{s_{xx}} \right\} MSE}$.

H.2.2 100(1- α)% confidence interval for β_1 : $\hat{\beta}_1 \mp t_{\alpha/2} \sqrt{MSE/s_{xx}}$

H.2.3 Testing the hypothesis $H_0 : \beta_1 = c$: $t = \frac{\hat{\beta}_1 - c}{\sqrt{MSE/s_{xx}}}$ (p364)

H.3 Confidence Interval

H.3.1 Confidence Interval for Mean for a given x : $\hat{y} \pm t_{\alpha/2} \sqrt{\left\{ \frac{1}{n} + \frac{(x - \bar{x})^2}{s_{xx}} \right\} MSE}$,

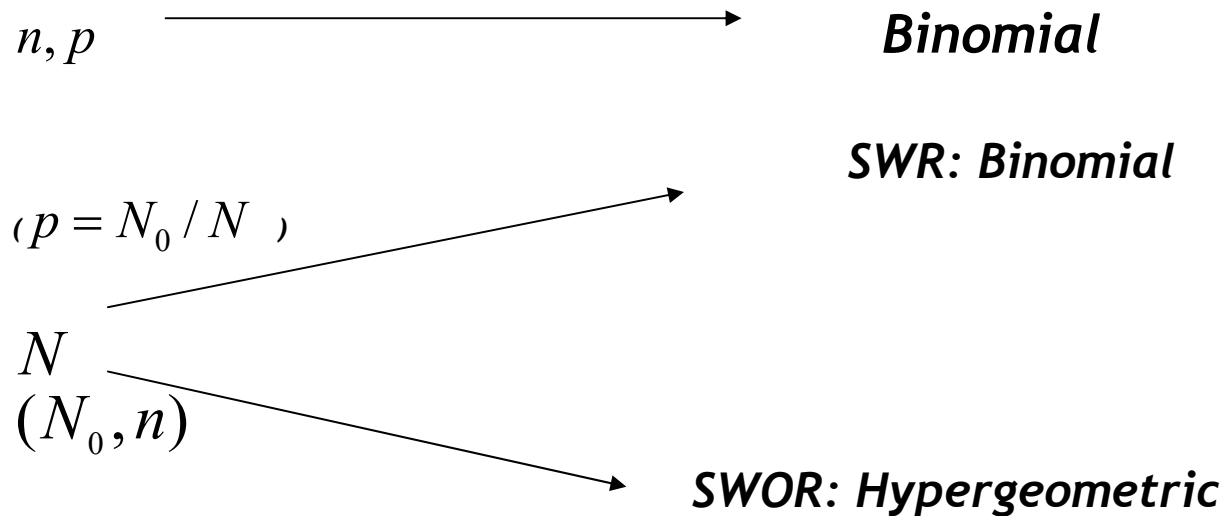
H.3.2 Prediction Interval for an Individual Y Given x :

$\hat{y} \pm t_{\alpha/2} \sqrt{\left\{ 1 + \frac{1}{n} + \frac{(x - \bar{x})^2}{s_{xx}} \right\} MSE}$ (p368)

Table for Estimation or Tests of Hypotheses on Mean or Difference in Means (In this table sample is always drawn randomly, and two samples considered are always independent)

| σ_i^2 | n_i | | Y | CI | z/t | TEST | page |
|---------------------------------|-----------------------|---|--------------------------------------|-------|-------------|-------|------|
| K N O W N | | 1 | Normal | F.1.0 | z | G.1.0 | 235 |
| | Large | 1 | nonnormal | F.1.1 | $\approx z$ | G1.1 | 235 |
| | | 2 | Normal | F.2.0 | z | G.2.0 | 247 |
| | Large | 2 | nonnormal | F.2.1 | $\approx z$ | G.2.1 | 247 |
| U N K N O W N | | 1 | Normal | F.1.3 | t | G.1.3 | 239 |
| | Large | 1 | Nonnormal | F.1.2 | $\approx z$ | G.1.2 | 239 |
| | | 2 | Normal ($\sigma_1^2 = \sigma_2^2$) | F.2.3 | t | G.2.3 | 249 |
| | Small ($n_1 = n_2$) | 2 | Normal | F.2.4 | $\approx t$ | G.2.4 | 251 |
| | | 2 | Normal | F.2.5 | $\approx t$ | G.2.5 | 251 |
| | Large | 2 | Nonnormal | F.2.2 | $\approx z$ | G.2.2 | 251 |

Distribution Guide



- **First success in a trial: Geometric**
- **Second/ third/.. success in a trail: Negative Binomial**
- **Mean & discrete variable: Poisson (# accidents occurs in a given time interval)**
- **Time between accidents: Exponential**

Lifetime : Exponential

mean = β , *parameter* = $\lambda = 1 / \beta$

